

# Bayesian and frequentist tests of sign equality and other nonlinear inequalities

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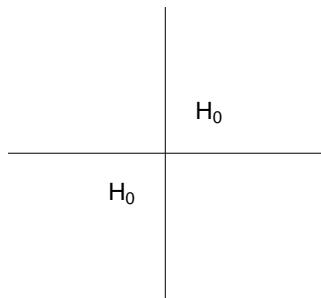
# Outline

- 1 Tests of sign equality
- 2 One dimension
- 3 General result
- 4 Unbiased test of sign equality
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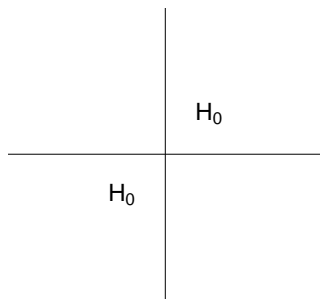
## Setup (1/2)

- Unknown:  $\theta_1 \in \mathbb{R}, \theta_2 \in \mathbb{R}$
- $H_0 : \theta_1\theta_2 \geq 0, H_1 : \theta_1\theta_2 < 0$



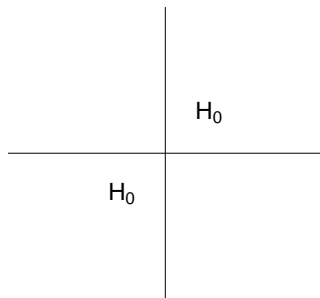
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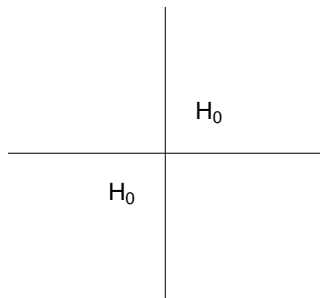
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- Ex (IV):  $y = x\beta + u$ ,  $x = z\pi + v$ ;  
 $y = z\pi\beta + \eta$ ,  $\theta_1 = \pi$ ,  $\theta_2 = \pi\beta$ ;  
 $H_0 : \beta \geq 0 \Leftrightarrow H_0 : \theta_1\theta_2 \geq 0$  since  
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 $\pi^2 \geq 0$
- Ex: stability over time,  $\theta_t$ ,  
 $t \in \{1, 2\}$



## Setup (2/2)

- Local-to-zero asy:  $\theta_1 = d_1/\sqrt{n}$ ,  $\theta_2 = d_2/\sqrt{n}$ ,

$$\sqrt{n} \begin{pmatrix} \hat{\theta}_1 \\ \hat{\theta}_2 \end{pmatrix} \xrightarrow{d} N \left( \begin{pmatrix} d_1 \\ d_2 \end{pmatrix}, V \right), \quad \hat{V} \xrightarrow{p} V.$$



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- Limit experiment: single draw of

$$\begin{pmatrix} \hat{t}_1 \\ \hat{t}_2 \end{pmatrix} \sim N \left( \begin{pmatrix} d_1 \\ d_2 \end{pmatrix}, \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix} \right),$$

known  $\rho$ , unknown  $d_1, d_2$ ;  $\hat{t} = \sqrt{n}\hat{\theta}/\hat{\sigma}$ .

# Likelihood ratio (LR) test: derivation

- Let  $\rho = 0$  for now; let  $\alpha = 0.05$
- LR: Euclidean distance from  $H_0$  (iso-LR: concentric circles)

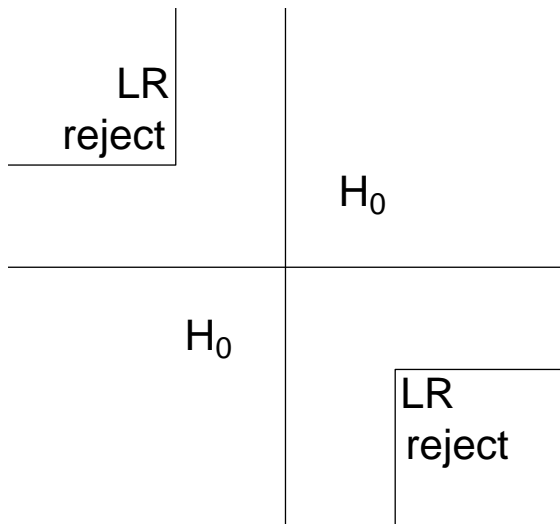
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- Very conservative near origin

## LR test: rejection region



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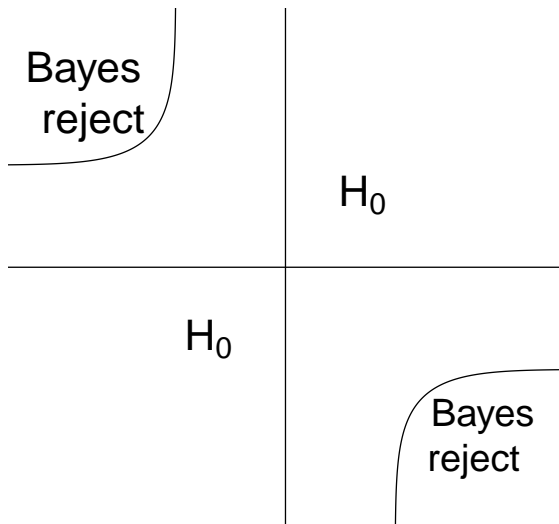
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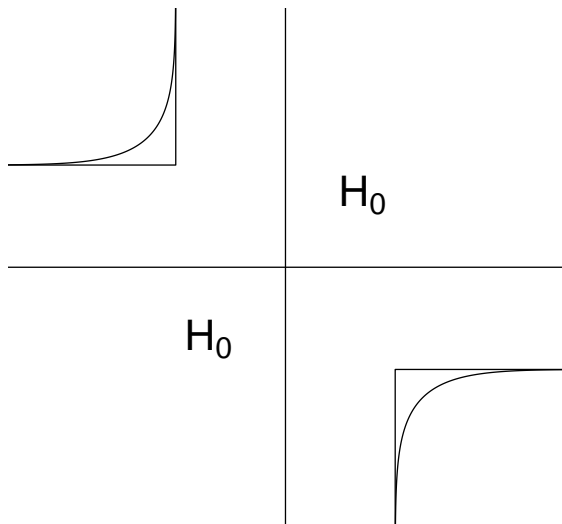
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- If  $P(H_0) = \alpha$ , then  $\mathbb{E}(L \mid \text{rej}) = \mathbb{E}(L \mid \text{acc}) = \alpha(1 - \alpha)$
- If  $P(H_0) < \alpha$ , should reject; if  $P(H_0) > \alpha$ , accept

## Bayesian test: rejection region



## LR vs. Bayesian rejection regions



# LR vs. Bayesian rejection regions

- I was surprised (just me?)
- Kline (2011):  $H_0 : \mu \geq 0$ ,  $\mu \in \mathbb{R}^k$ , can have  $P(H_0) \approx 0$  and frequentist  $p$ -value  $\approx 1$
- Moon and Schorfheide (2012, Cor. 1): for set-identified parameter, Bayesian credible interval is strict subset of frequentist CI
- Bayes rule vs. minimax risk



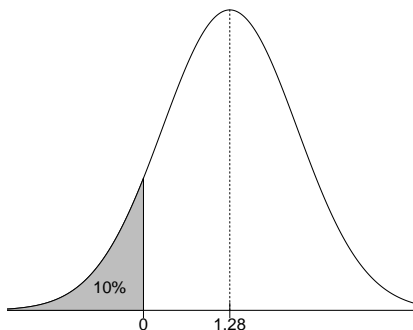
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# 1D: standard one-sided testing

- Sampling dist:  $X \sim N(\mu, 1)$ ; posterior:  $\mu \sim N(X, 1)$
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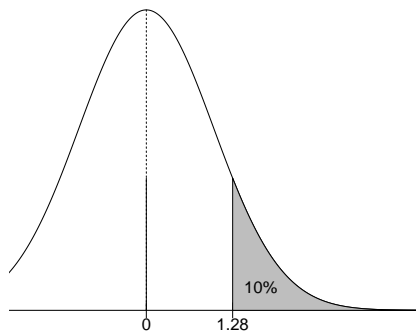
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- Bayesian:  $P(H_0 | X = 1.28) = \alpha$ ; reject if  $X > 1.28$ , otherwise accept



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- Frequentist:  $P(X > 1.28 \mid \mu = 0) = \alpha$ ; reject if  $X > 1.28$ , otherwise accept



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- “Unbiased”: rejection probability (RP) given any  $\mu \leq 0$  is  $\leq$  RP given any  $\mu > 0$
- Minimax risk (expected loss) decision rule:
  - Same loss as before:  $1 - \alpha$  for type I,  $\alpha$  for II
  - Let  $RP(\mu)$  be RP given  $\mu$ , continuous in  $\mu$
  - Given  $\mu \leq 0$ , risk is  $(1 - \alpha)RP(\mu)$
  - Given  $\mu > 0$ , risk is  $\alpha(1 - RP(\mu))$
  - $RP(\mu = 0) = RP(\mu \rightarrow 0+)$ , so min max risk by  $RP(0) = \alpha$

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- $X \sim N(\mu, 1), \mu \sim N(X, 1)$
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  - $P(|X| > 1.64 | \mu = 0) = \alpha \implies$  rejecting if  $|X| > 1.64$  is valid freq test



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- Small  $h$ : similar
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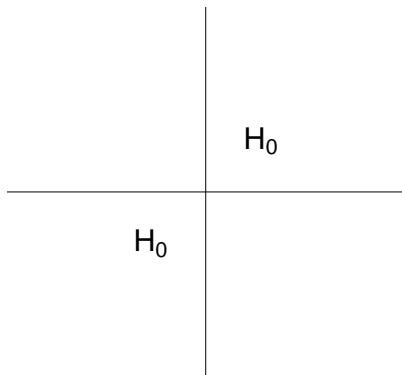
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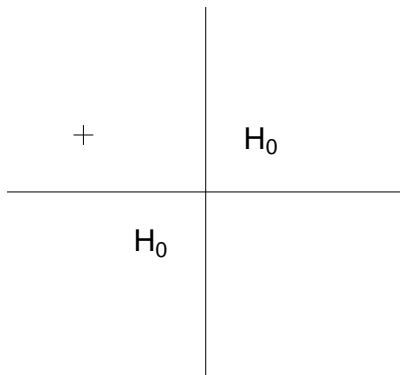
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- $H_0 : d_1 d_2 \geq 0$ ,  $H_1 : d_1 d_2 < 0$
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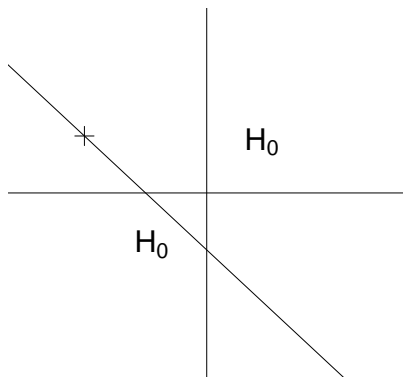
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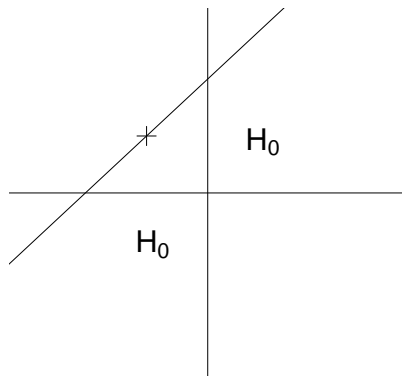
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# General result

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- (Corollary 6) Linear inequalities produce convex  $\Theta_0$ , so frequentist testing is more conservative; ex:  $H_0 : \theta \geq 0$

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# Unbiased moment inequality tests

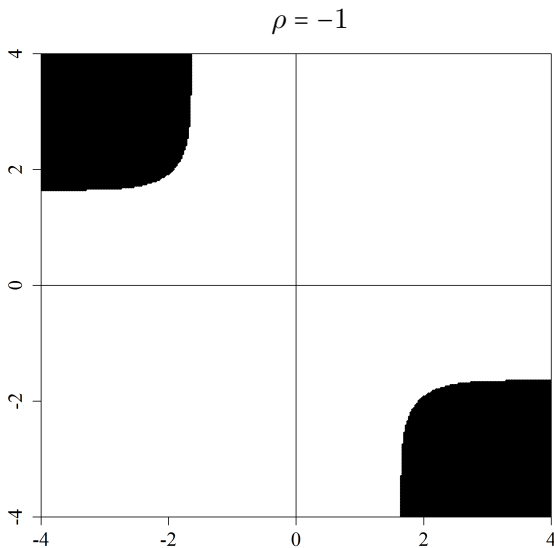
- Andrews (2012): “Similar-on-the-boundary tests for moment inequalities exist, but have poor power” for  $H_0 : \mu \geq 0$  (vector)

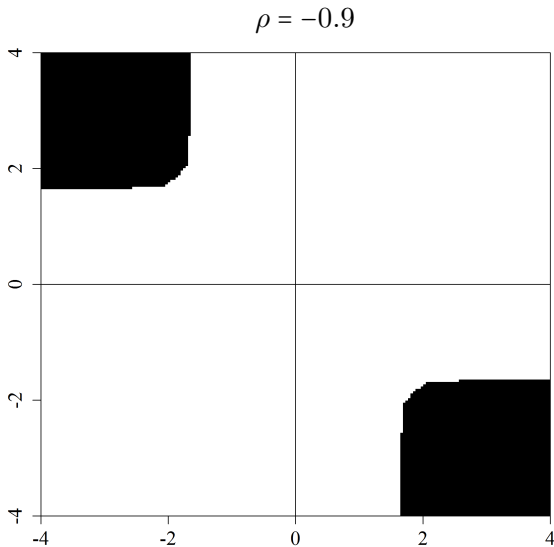
# Unbiased moment inequality tests

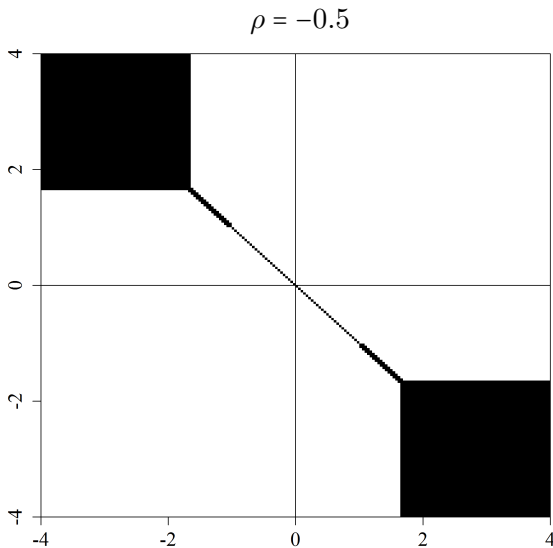
- Andrews (2012): “Similar-on-the-boundary tests for moment inequalities exist, but have poor power” for  $H_0 : \mu \geq 0$  (vector)
- Argument doesn't apply here since  $\{\theta : \theta_1\theta_2 = 0\}$  is entirely on boundary of  $H_0$  for sign equality
- But: applies to sign equality in higher dimensions

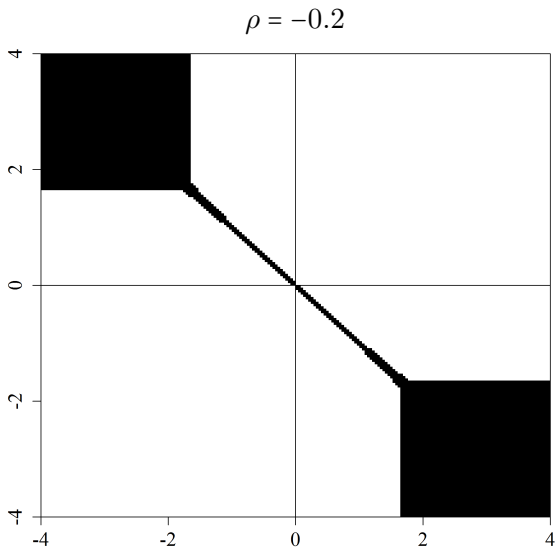
# Unbiased sign equality test

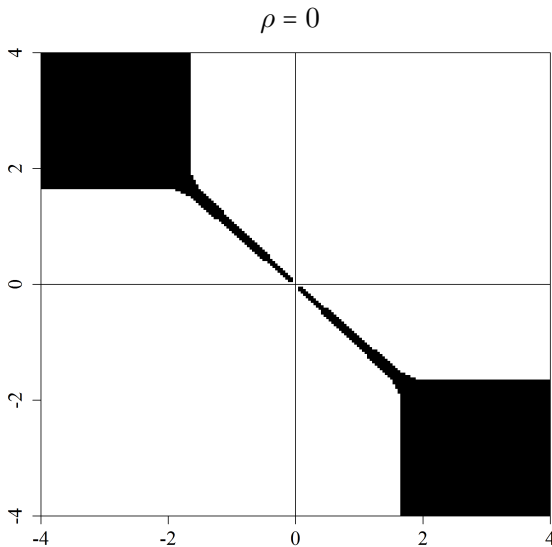
- How to derive? Chiburis (2008), Moreira and Moreira (2013): frame as linear program, max weighted avg power (WAP) s.t. (approx) unbiasedness. See also Berger (1989, §6) for  $\rho = 0$
- Pro: unbiased; minimax risk
- Pro: power dominates LR over  $H_1$  (but risk fn doesn't dominate over  $H_0$ )
- Con: computation; LR and Bayesian tests much simpler/faster
- Con: unbiased test lacks a type of “rejection monotonicity,” and (for larger  $\rho > 0$ ) can reject even if  $\hat{t}$  satisfies  $H_0$ ; see Perlman and Wu (1999) for more detailed criticism

Rejection regions for various  $\rho$ ,  $\alpha = 0.05$ 

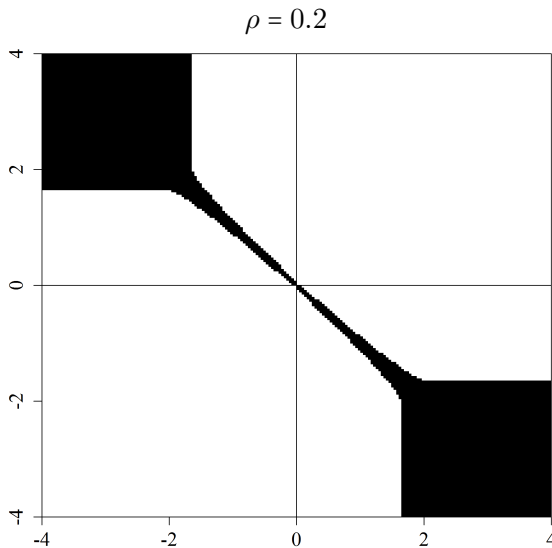
Rejection regions for various  $\rho$ ,  $\alpha = 0.05$ 

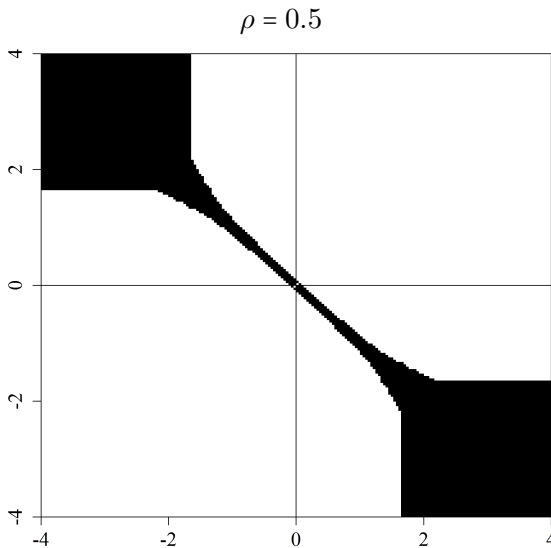
Rejection regions for various  $\rho$ ,  $\alpha = 0.05$ 

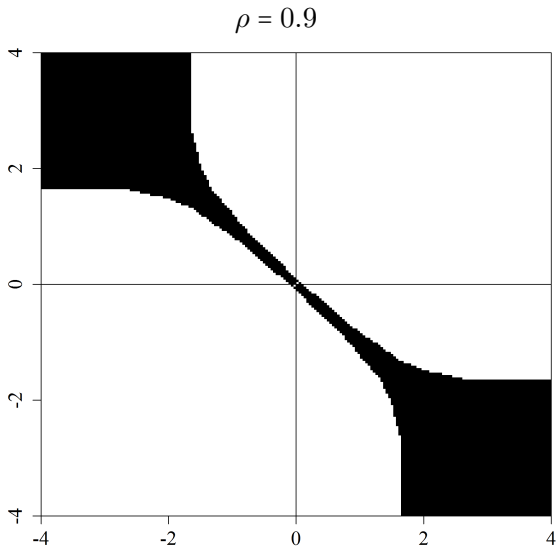
Rejection regions for various  $\rho$ ,  $\alpha = 0.05$ 

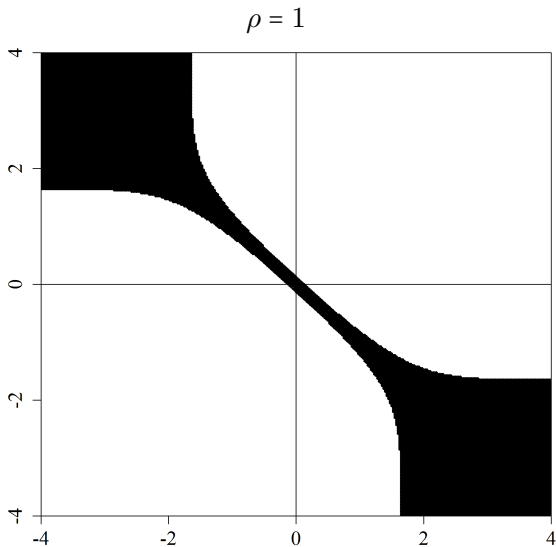
Rejection regions for various  $\rho$ ,  $\alpha = 0.05$ 



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- 1 Tests of sign equality
- 2 One dimension
- 3 General result
- 4 Unbiased test of sign equality
- 5 Conclusion

# Recap

- Frequentist testing is “more conservative” than Bayesian testing with convex  $\Theta_0$ , given equivalent asymptotic normal sampling and posterior distributions
- Linear inequalities always define a convex  $\Theta_0$
- Sign equality  $\Theta_0$  is non-convex; which is more conservative depends on  $\rho$
- Unlike for  $H_0 : \theta \geq 0$ , unbiased tests with good power exist for  $H_0 : \theta_1\theta_2 \geq 0$  (but not higher dimensions)
- But, unbiasedness comes at a cost

# Bayesian and frequentist tests of sign equality and other nonlinear inequalities

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# References I

- Andrews, D. W. K. (2012). Similar-on-the-boundary tests for moment inequalities exist, but have poor power. Discussion paper 1815R, Cowles Foundation for Research in Economics at Yale University, New Haven, CT.
- Berger, R. L. (1989). Uniformly more powerful tests for hypotheses concerning linear inequalities and normal means. *Journal of the American Statistical Association*, 84(405):192–199.
- Chiburis, R. C. (2008). Approximately most powerful tests for moment inequalities. Job market paper, Princeton University, Princeton, NJ.
- Kline, B. (2011). The Bayesian and frequentist approaches to testing a one-sided hypothesis about a multivariate mean. *Journal of Statistical Planning and Inference*, 141(9):3131–3141.



## References II

- Moon, H. R. and Schorfheide, F. (2012). Bayesian and frequentist inference in partially identified models. *Econometrica*, 80(2):755–782.
- Moreira, H. and Moreira, M. J. (2013). Contributions to the theory of optimal tests. *Ensaaios Econômicos 747*, FGV/EPGE, Rio de Janeiro.
- Perlman, M. D. and Wu, L. (1999). The emperor's new tests. *Statistical Science*, 14(4):355–369.